Rainfall manipulation experiments as simulated by terrestrial biosphere models: Where do we stand?

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See also the Commentary on this article by Xue Feng, 26, 3190–3192
Abstract
Changes in rainfall amounts and patterns have been observed and are expected to continue in the near future with potentially significant ecological and societal consequences. Modelling vegetation responses to changes in rainfall is thus crucial to project water and carbon cycles in the future. In this study, we present the results of a new model-data intercomparison project, where we tested the ability of 10 terrestrial biosphere models to reproduce the observed sensitivity of ecosystem productivity to rainfall changes at 10 sites across the globe, in nine of which, rainfall exclusion and/or irrigation experiments had been performed. The key results are as follows: (a) Inter-model variation is generally large and model agreement varies with timescales. In severely water-limited sites, models only agree on the interannual variability of evapotranspiration and to a smaller extent on gross primary productivity. In more mesic sites, model agreement for both water and carbon fluxes is typically higher on fine (daily–monthly) timescales and reduces on longer (seasonal–annual) scales. (b) Models on average overestimate the relationship between ecosystem productivity and mean rainfall amounts across sites (in space) and have a low capacity in reproducing the temporal (interannual) sensitivity of vegetation productivity to annual rainfall at a given site, even though observation uncertainty is comparable to inter-model variability. (c) Most models reproduced the sign of the observed patterns in productivity changes in rainfall manipulation experiments but had a low capacity in reproducing the observed magnitude of productivity changes. Models better reproduced the observed productivity responses due to rainfall exclusion than addition. (d) All models attribute ecosystem productivity changes to the intensity of vegetation stress and peak leaf area, whereas the impact of the change in growing season length is negligible. The relative contribution of the peak leaf area and vegetation stress intensity was highly variable among models.

KEYWORDS
drought, irrigation, rainfall manipulation experiment, terrestrial biosphere models

1 | INTRODUCTION
Understanding the impact of rainfall changes on ecosystem functioning and vegetation dynamics is crucial for accurately predicting the responses of vegetation structure, composition and dynamics under present or future conditions. Changes in both rainfall intensity and variability have been measured in the last decades (IPCC, 2013; Trenberth, 2011). Changes in precipitation extremes have also been observed (Alexander et al., 2006) and according to climate model projections, such changes will intensify as we progress through the 21st century (IPCC, 2012; Knutti & Sedláček, 2013).

Changes in rainfall can affect energy and carbon fluxes at the land surface (Green et al., 2017). Rainfall changes modify soil water dynamics, alter plant water status and consequently the terrestrial biogeochemical cycles (Allan et al., 2014; Heisler-White, Knapp, & Kelly, 2008) through changes in plant productivity or plant mortality (Allen, Breshears, & McDowell, 2015). The importance of plant water limitation has been highlighted by the fact that semi-arid regions, which typically experience drought, control part of the global interannual variability of the terrestrial carbon sink (Ahlström et al., 2015), with an increasing sensitivity during the last decades (Poulter et al., 2014). The importance of water limitation on carbon fluxes will likely increase soon, since terrestrial vegetation is thought to operate close to its critical hydraulic thresholds across a wide range of ecosystems (Choat et al., 2012), even though the full implications of this result are still debated (Klein, Yakir, Buchmann, & Grünzweig, 2014; Körner, 2019). As a direct consequence, minor changes in plant water availability worldwide can lead to significant impacts on the terrestrial carbon sink (Allen et al., 2010; Frank et al., 2015; Green et al., 2019; Humphrey et al., 2018; Reichstein et al., 2013; Zhao & Running, 2010).

To understand the ecosystem responses to changes in rainfall amounts and patterns at the local scale, rainfall manipulation experiments have been conducted. Typically, such experiments change the
overall rainfall amount by exclusion (Estiarte et al., 2016; Limousin et al., 2009; Martin-StPaul et al., 2013) or irrigation (Collins et al., 2012) and responses are commonly quantified by the changes in aboveground net primary production (ANPP). In some experiments such as the Amazon rainfall exclusion experiment (Nepstad, Tohver, Ray, Moutinho, & Cardinot, 2007), additional detailed data quantifying the changes in forest structure and composition have been obtained. There are a small number of experiments where the structure of rainfall pulses is modified (e.g. Faye, Kaufman, Nippert, Carlisle, & Harper, 2008; Heisler-White et al., 2008; Vicca et al., 2014). Rainfall manipulation experiments have been conducted in a range of ecosystems, spanning from semi-arid shrublands (Baez, Collins, Pockman, Johnson, & Small, 2013) to temperate (Hanson & Wullschleger, 2003) and tropical forests (Fisher et al., 2007; Nepstad et al., 2007), even though most of the experiments have focused on grasslands or low stature vegetation due to the difficulties in setting up experiments. Those experiments have identified a strong correlation between rainfall changes and vegetation productivity (e.g. Heisler-White, Blair, Kelly, Harmoney, & Knapp, 2009; Stuart-Haëntjens et al., 2018), phenology (e.g. Peruelas et al., 2004), plant community structure (e.g. Miranda, Armas, Padilla, & Pugnaire, 2011; Zhang et al., 2019) and belowground carbon dynamics (e.g. Hagedorn et al., 2016; Hasibeder, Fuchsleuger, Richter, & Bahn, 2014; Vicca et al., 2014). Despite the important findings derived from these field experiments, these studies have strong spatial and temporal limitations; they reported only few variables and it is challenging to extrapolate information beyond the specific design of the experiment. Extrapolation and mechanistic understanding related to vegetation responses to changes in precipitation can be better achieved by combining model and data-driven approaches (e.g. Kayler et al., 2015).

Modelling vegetation responses to changes in water availability is a challenging task (Xu, McDowell, Sevanto, & Fisher, 2013). Despite strong evidence that modelling responses to drought is a significant factor affecting terrestrial carbon dynamics (Trugman, Medvigy, Mankin, & Anderegg, 2018), a commonly accepted parameterization of water limitation does not exist (Egea, Verhoef, & Vidale, 2011; Fatichi, Pappas, & Ivanov, 2016; Hu et al., 2018; Medlyn, Kauwe, & Duursma, 2016; Zhou, Duursma, Medlyn, Kelly, & Prentice, 2013). Plant water stress simulated in terrestrial biosphere models can affect various processes but is commonly a function of either volumetric soil water content (e.g. Clark et al., 2011) or soil water potential (e.g. Fatichi, Ivanov, & Caporali, 2012; Lawrence et al., 2019; Manzoni, Vico, Porporato, & Katul, 2013), integrated over the root zone. Examples of how water limitation affects plant functions include a decline in stomatal conductance affecting photosynthesis (De Kauwe, Kala, et al., 2015; De Kauwe, Zhou, et al., 2015; Egea et al., 2011; Fatichi et al., 2012), changes in the photosynthetic parameters $V_{c,max}$ and $J_{max}$ (e.g. Krinner et al., 2005) and/or accelerated senescence of plant tissues, especially leaves (Thurner et al., 2017) leading to drought-induced deciduousness. Recently, significant efforts have been made to include more detailed plant hydraulics, to better describe water flow within the soil-plant-water continuum (Bonan, Williams, Fisher, & Oleson, 2014; Eller et al., 2018; Kennedy et al., 2019; Lawrence et al., 2019; Mirfenderesgi et al., 2016) and to include dynamics of non-structural carbohydrates to simulate consequences of water stress for carbon allocation and carbon starvation (reviewed in Fatichi, Pappas, Zscheischler, & Leuzinger, 2019).

A large discrepancy of predicted model responses has direct consequences for the uncertainties related to the fate of terrestrial carbon under a changing climate (Ahlström et al., 2015; Humphrey et al., 2018; Zscheischler, Michalak, et al., 2014). This is the case because the terrestrial vegetation and thus the terrestrial land carbon sink introduce the largest uncertainties of the global carbon cycle (Le Quéré et al., 2018). In this context, large epistemic model uncertainties can have considerable impacts on our ability to forecast the growth rate of atmospheric CO$_2$. Additionally, vegetation responses to water stress can influence land-atmosphere coupling (Gentine et al., 2019; Koster, 2004; Lemordant, Gentine, Stéfanon, Drobinski, & Fatichi, 2016; Seneviratne et al., 2013), since vegetation cover and canopy conductance affect land surface energy balance. This will have a large impact on our skill to model the coupled hydrological, plant physiological and meteorological processes and thus robustly projecting climate change (Miralles, Gentine, Seneviratne, & Teuling, 2018).

To reduce this source of epistemic uncertainty and understand the reasons for model disagreement, a detailed comparison between the responses of different modelling schemes with respect to plant water availability is essential. Rainfall manipulation experiments assessing vegetation responses to water limitation are particularly useful in this regard. Arguably, this is an extremely important test to evaluate the structure and parameter values of a model and its capability to reproduce responses to environmental changes. A model should be able to reproduce the observed dynamics under control and manipulated conditions in order to be considered robust, especially for climate change simulations (Medlyn et al., 2015). Despite the importance of this comparison, there are only a few examples that have compared terrestrial biosphere models and global change manipulation experiments (De Kauwe et al., 2013, 2017; Fatichi & Leuzinger, 2013; Medlyn et al., 2015; Powell et al., 2013; Zaehle et al., 2014). Recently, Wu et al. (2018) compared 14 models under different idealized rainfall scenarios for three grassland experiments sites and showed a fair reproduction of spatial sensitivities of ANPP to rainfall but large differences in the modelled asymmetric response of ANPP to interannual, that is temporal rainfall variability at a given site. Wu et al. (2018) were not able to evaluate the modelled responses with respect to actual experiments because they used idealized rainfall changes that did not exactly mimic the site treatments. In this study we perform such an evaluation. We make use of 10 sites with diverse climates and biomes, where multiyear rainfall manipulation experiments took place to evaluate 10 terrestrial biosphere models, representing an unprecedented data-model intercomparison effort focused on ecosystem responses to water limitation. This data-model intercomparison will address the following questions: (a) Can models reproduce the observed responses to precipitation variability at rainfall manipulation sites? (b) Do models accurately reproduce the spatial (across-sites) and temporal (within-site) dependence of vegetation productivity to precipitation? (c) Which are the underlying reasons for model disagreement? Answering those questions will provide insights on the robustness of Earth System model projections with respect to the global carbon cycle.
2 | DATA AND METHODS

2.1 | Sites

Ten different sites with contrasted climates and biomes and sufficiently long records were considered here. For all analyses presented in this study, the sites are termed: Lahav, Matta, SGS, Prades, Garraf, Konza (AmeriFlux ID: US-Kon), Puèchabon (FluxNet ID: FR-Pue), Brandbjerg, Walker Branch (Walker Branch; AmeriFlux ID: US-WBBW) and Stubai (Table 1). The sites are in ascending order in terms of wetness index (WI) defined as the average ratio of annual precipitation to annual potential evapotranspiration (ET) during the study period. For our analysis the sites are split in three wetness categories (WI < 0.4 [Lahav, Matta, SGS]; 0.4 ≤ WI < 1 [Prades, Garraf, Konza, Puèchabon]; and WI ≥ 1 [Brandbjerg, Walker Branch, Stubai]).

The sites are in the United States (Konza, SGS, Walker Branch), Israel (Lahav, Matta), Spain (Garraf), France (Prades, Puèchabon), Austria (Stubai) and Denmark (Brandbjerg) and span a precipitation gradient from 253 to 1,440 mm/year and include grasslands shrublands and forested ecosystems (Table 1). In eight for the sites rainfall exclusion experiments were carried out, and in four of the sites, irrigation experiments were carried out. The experiment duration considered in this study was from 5 up to 32 years. The average experiment duration was 13.3 years.

For all sites, ANPP estimates were recorded for most of the experimental years derived by either biomass harvesting (grasslands) or biomass increase estimates derived from allometric relations and simultaneous observations of stem diameter, leaf area changes, plus litterfall (e.g. shrublands and forests). Leaf area index (LAI) was quantified using the MODIS (MCD15A2H v006) estimate of the pixel containing each site. MODIS data were interpreted with caution as they are an indirect measurement, valid at typically larger scales, and prone to large uncertainties. For three sites, Konza, Puèchabon and WB, ET and gross primary productivity (GPP) were obtained at the half hourly scale by the Fluxnet2015 database and aggregated to the daily scale.

2.2 | Participating models and simulation protocol

For all sites, we conducted simulations using 10 terrestrial biosphere models: CABLE r54482.0 (Wang et al., 2011), DLEM v2.0 (Tian et al., 2010), JULES v5.2 (Clark et al., 2011), JSBACH v3.2 (Kaminski et al., 2013; Mauritsen et al., 2019), LPX v1.4 (Lienert & Joos, 2018), ORCHIDEE rev5150 (Krinner et al., 2005), ORCHIDEE MICT rev5308 (Guimberteau et al., 2018), ORCHIDEE CNP rev4520 (Goll et al., 2017), T&C v1.0 (Faticchi et al., 2012; Paschalis, Katul, Faticchi, Palmroth, & Way, 2017) and TECO v2.0 (Huang et al., 2017). All models include a land surface scheme, a hydrological component and a dynamic vegetation module. Soil moisture dynamics are simulated in multiple vertical layers by either solving the 1D Richards equation or simplified hydrological 'bucket-type' models. Some models can simulate vegetation succession; however, this feature was disabled in the current study. Five models included nutrient dynamics. CABLE, DLEM, JSBACH and LPX simulated nitrogen and ORCHIDEE CNP nitrogen/phosphorus cycles. Hydrological and biogeochemical processes are simulated with a variable degree of complexity (for a detailed model description see the supplementary material of Wu et al., 2018). As there is no commonly accepted way to simulate water limitation, each model has adopted significantly different approaches (Medlyn, De Kauwe, Zaehle, et al., 2016). Water stress in all models but T&C is a function of an average root zone soil moisture; and in T&C, water stress is a function of the integrated root zone soil water potential. Specifically, models alter either photosynthetic rates (T&C, JULES, TECO), the maximum rate of carboxylation \( V_{\text{cmax}} \) (ORCHIDEE, ORC MICT, ORC CNP), stomatal conductance (JSBACH, DLEM) or a combination of all such parameters (CABLE), based on plant water availability. LPX uses a supply and demand-driven approach to water limitation. If water demand exceeds supply, photosynthesis is downregulated until they match. None of the models simulates plant hydraulics and thus xylem cavitation in response to water stress.

For each site, we conducted a control simulation corresponding to the observed climate without manipulation, and simulations representative of each rainfall manipulation experiment (rainfall exclusion and/or irrigation) with the same timing and magnitude of water input as in the real experiment. For all experiments the common data distributed to all modelling groups included hourly values of incoming shortwave and longwave radiation, vapour pressure deficit, air temperature, wind speed, atmospheric pressure and ambient CO\(_2\) concentration. Model set-up was performed by each modelling group separately based on common information for each site that included, apart from the meteorological input, species composition, vegetation cover, soil and root depth, and soil textural properties. Each modelling group translated independently this information into model-specific parameters. Dependent on the model, species composition and vegetation cover were used to either choose between prescribed plant functional types (PFTs) or plant-specific model parameters. Soil and root depth were used by all modelling groups to set-up the simulation domain, and the vertical discretization of the simulation was decided by each modelling group independently. Soil textural properties were used to select soil hydraulic properties. All information concerning the simulation set-up of each model and the common site properties provided to all modelling groups can be found at a free access data repository (see Data Availability Statement). Reported model outputs included GPP, NPP and ANPP, ET and its partition in evaporation (soil evaporation plus evaporation from interception) and transpiration, respectively, soil moisture, LAI and biomass carbon pool (below and aboveground) dynamics. Some models additionally reported the water stress factor (\( \beta \)) used in the model. \( \beta \) is a model parameter that quantifies the effects of plant physiological stress due to limitations in soil water availability. \( \beta \) is not identical between models and the description of the \( \beta \) for each model can be found at the supplementary material of Wu et al. (2018). Initial conditions for all simulations were obtained after a spin-up period long enough to equilibrate the biogeochemical cycles.
<table>
<thead>
<tr>
<th>Site</th>
<th>Lon/Lat</th>
<th>Annual T (°C)</th>
<th>Annual P (mm)</th>
<th>Altitude (m)</th>
<th>Species</th>
<th>Soil type</th>
<th>Drought treatment</th>
<th>Irrigation treatment</th>
<th>Years</th>
<th>Key references</th>
</tr>
</thead>
<tbody>
<tr>
<td>Lahav</td>
<td>34.9/</td>
<td>19.1</td>
<td>253</td>
<td>0.19</td>
<td>Annual grasses and shrubs, mostly Sarcopoterium spinosum</td>
<td>22.6% Sand, 39.7% silt and 37.7% clay</td>
<td>−30% rainfall for the entire year</td>
<td>+30% rainfall for the entire year</td>
<td>2002–2014</td>
<td>Tielbörger et al. (2014)</td>
</tr>
<tr>
<td>Matta</td>
<td>35.07/</td>
<td>17.94</td>
<td>498</td>
<td>0.33</td>
<td>Similar to Lahav</td>
<td>19% Sand, 29.2% silt and 51.8% clay</td>
<td>−30% rainfall for the entire year</td>
<td>+30% rainfall for the entire year</td>
<td>2002–2014</td>
<td>Tielbörger et al. (2014)</td>
</tr>
<tr>
<td>SGS</td>
<td>−104.75/</td>
<td>8.4</td>
<td>304</td>
<td>0.35</td>
<td>C4 grasses, primarily (Bouteloua gracilis H.B.K.) Lag. Ex Steud., Buchloe dactyloides (Nutt) Engelm., mixed with varying amounts of C3 grasses, cactus, shrubs and forb.</td>
<td>14% Sand, 58% silt and 28% clay</td>
<td>None</td>
<td>None</td>
<td>1986–2009</td>
<td>Heisler-White et al. (2009)</td>
</tr>
<tr>
<td>Prades</td>
<td>0.91/</td>
<td>11.43</td>
<td>522</td>
<td>0.4</td>
<td>Mixed composition of Quercus ilex L., Phillyrea latifolia L., Arbutus unedo L., Erica arborea L., Juniperus oxycedrus L., Cistus albidus L, Sorbus torminalis (L.) Crantz and Acer monspessulanum L.</td>
<td>48% Sand, 32% silt and 20% clay</td>
<td>−20% rainfall for the entire year</td>
<td>None</td>
<td>1999–2012</td>
<td>Ogaya and Peñuelas (2007)</td>
</tr>
<tr>
<td>Garraf</td>
<td>1.82/</td>
<td>15.04</td>
<td>580</td>
<td>0.48</td>
<td>Erica multiflora, Globularia alypum</td>
<td>41% Sand, 39% silt and 18% clay</td>
<td>−50% in spring and fall</td>
<td>None</td>
<td>2000–2004</td>
<td>Beier et al. (2009)</td>
</tr>
<tr>
<td>Konza</td>
<td>−96.6/</td>
<td>12.8</td>
<td>830</td>
<td>0.7</td>
<td>Mixed C3 (Solidago canadensis, Aster ericoides, Salix missouriensis) C4 (Andropogon gerardii, Sorghastrum nutans, Panicum virgatum) Grassland</td>
<td>10% Sand and 35% clay</td>
<td>None</td>
<td>Irrigation +20% was provided at two sites termed lowland and upland</td>
<td>1982–2013</td>
<td>Collins et al. (2012)</td>
</tr>
<tr>
<td>Puèchabon</td>
<td>43.74/</td>
<td>13.8</td>
<td>969</td>
<td>0.87</td>
<td>Ouerstory (Q. ilex); Understory (Buxus sempervirines, P. latifolia, Pistacia terebinthus and J. oxycedrus)</td>
<td>26% Sand, 35% silt and 39% clay</td>
<td>−30% throughfall exclusion for the entire year</td>
<td>None</td>
<td>2004–2013</td>
<td>Limousin et al. (2009)</td>
</tr>
<tr>
<td>Brandbjerg</td>
<td>11.97/</td>
<td>9.59</td>
<td>757</td>
<td>1.1</td>
<td>70% grasses (mostly Deschampsia flexuosa); 30% dwarf shrubs (Calluna vulgaris)</td>
<td>88%–95% Sand, 2%–9% silt, and 1%–2% clay</td>
<td>Rainfall exclusion for 4–6 weeks during spring and summer</td>
<td>None</td>
<td>2007–2012</td>
<td>Kongstad et al. (2012)</td>
</tr>
<tr>
<td>WB</td>
<td>−84.29/</td>
<td>14.7</td>
<td>1,440</td>
<td>1.1</td>
<td>Mixed composition of Quercus spp; Quercus prinus L., Quercus alba L., Quercus rubra L., Acer rubrum L., Acer saccharum, Liriiodendron tulipifera L., Nyssa sylvatica Marsh, and Oxydendrum arboreum (L.)</td>
<td>28% Sand, 60% silt and 12% clay</td>
<td>−30% throughfall exclusion for the entire year</td>
<td>+33% rainfall for the entire year</td>
<td>1995–2005</td>
<td>Hanson et al. (2004)</td>
</tr>
<tr>
<td>Stubai</td>
<td>11.32/</td>
<td>6.8</td>
<td>1,382</td>
<td>1.7</td>
<td>C3 Grassland (Agrostis capillaris, Festuca rubra, Ranunculus montanus, Trifolium pratense, Trifolium repens)</td>
<td>42.2% Sand, 47% silt and 10.8% clay</td>
<td>Rainfall exclusion for 8 weeks of summer rainfall</td>
<td>None</td>
<td>2009–2013</td>
<td>Fuchslueger, Bahn, Fritz, Hasibeder, and Richter (2014), Hasibeder et al. (2014)</td>
</tr>
</tbody>
</table>

Abbreviation: WI, wetness index.
2.3 | Statistical analyses

2.3.1 | Data-model comparison

First, we compare the models’ ability to accurately reproduce the relationship between ANPP and precipitation ($P$) across sites (i.e. spatial dependence) and within each site (i.e. temporal dependence) at the annual scale. At all sites, observations of ANPP were based on biomass estimates (e.g. using aboveground biomass harvesting for grasslands and a carbon budget approach for forested sites combining litterfall observations with allometric equation for aboveground biomass growth) rather than carbon fluxes, therefore discrepancy between observed and modelled ANPP is expected (detailed bias quantification are reported in the Supporting Information).

Model skill in reproducing the spatial dependence of ANPP to $P$ was quantified as the root mean squared error (RMSE) and the coefficient of determination ($R^2$) between the modelled and observed annual ANPP, averaged over the entire period, across sites for the control case. Model performance in capturing the magnitude of interannual variability of ANPP was assessed by comparing the standard deviation ($\sigma$) of annual ANPP between models and observations for all sites. Model skill with respect to single-site interannual dependence of ANPP to $P$ was quantified using an estimate of the sensitivity of annual ANPP to annual $P$. Specifically, we fitted a linear model $\text{ANPP} = a_0 + a_1 P + a_2 T$, where $P$ is annual precipitation and $T$ is annual temperature. To increase the sample size and robustness of the fit, precipitation from both the control and the rainfall manipulation experiments were used. Additional covariates such as vapour pressure deficit and radiation could not be added due to the small sample size, making the linear fit over constrained. Preliminary analyses (not reported here) showed that $P$ and $T$ were the most important covariates. Model skill was evaluated by estimating the differences between observed and simulated sensitivities of ANPP with respect to $P$ (i.e. $a_1 = \Delta \text{ANPP}/\Delta P$). Observation uncertainty of the sensitivity metric was quantified as the 90% confidence interval of the linear model fit.

For the control simulations, modelled ET and GPP were compared with eddy covariance high-frequency observations from Walker Branch, Puèchabon and Konza. In these three locations, flux-tower data were available in the proximity and with the same vegetation cover as the rainfall exclusion/addition experiment. Comparison at the daily scale was performed by means of Taylor diagrams (Taylor, 2001). The magnitude and seasonal pattern of the fluxes were also analysed (Figures S2–S4).

Responses due to rainfall manipulation were quantified at the annual scale using the response ratio for a variable $X$ (e.g. ANPP) defined as the ratio $RR = X_C^M / X_C^M$, where the subscript $M$ denotes manipulation, $C$ denotes the control scenario and $\bar{X}$ indicates the annual scale. In this study, we focused on the simulated RRs of ANPP and ecosystem water use efficiency (WUE) calculated at the annual scale as the ratio of annual GPP to annual actual ET. To quantify whether the simulated response ratios have a statistically significant different mean value from the observations, a two-sample $t$ test was performed for every model and the respective observed responses.

For the two-sample $t$ test, the sample size for each site is equal to the number of years in the observations and simulations. Response ratios were assumed normally distributed and independent at the annual scale. The test’s null hypothesis was that modelled and observed response ratios have the same mean. The analysis was also performed using the commonly used logarithm of RR yielding identical results, and thus not further shown here.

2.3.2 | Model agreement

Model agreement across timescales was quantified by estimating the Pearson correlation coefficient ($\rho$) between all pairs of models for ET and GPP at the daily, monthly and annual scale. In Figure S7, the analysis is expanded for a wider range of scales by estimating the wavelet coherence between all pairs of models for ET and GPP.

To quantify agreement with respect to modelled changes in ANPP and WUE due to rainfall alterations, a two-sample $t$ test for the response ratios of both ANPP for all model pairs was performed and presented in Tables S2 and S3.

To attribute the variability of ANPP to its causes we proceeded similar to De Kauwe et al. (2017) who found that the annual ANPP could be approximated by the product.

$$\text{ANPP} = A_0 \cdot \text{CUE} \cdot \text{GPP}_u \cdot \beta \cdot \text{LAI}_p \cdot \text{LAI}_r,$$

where $A_0$ is the aboveground fraction of carbon allocation, CUE is the carbon use efficiency, GPP$_u$ is the potential (unstressed) rate of GPP per unit of leaf area, $\beta$ is the annually averaged value of the water stress factor, LAI$_p$ is the peak LAI during the year and LAI$_r$ is the proxy of the growing season length, defined as the integral of LAI during the year divided by LAI$_r$. Considering that water stress and LAI dynamics determine most of the interannual variation of ANPP, assuming that $A_0$, CUE and GPP$_u$ vary less between treatments, then, the annual response ratio of ANPP can be estimated by the response ratios of $\beta$, LAI$_p$ and LAI$_r$.

$$\frac{\text{ANPP}_M^{\text{CUE}}}{\text{ANPP}_C^{\text{CUE}}} \approx \frac{\text{ANPP}_M^{\text{GPP}_u}}{\text{ANPP}_C^{\text{GPP}_u}} \cdot \frac{\text{ANPP}_M^{\beta}}{\text{ANPP}_C^{\beta}} \cdot \frac{\text{ANPP}_M^{\text{LAI}_p}}{\text{ANPP}_C^{\text{LAI}_p}} \cdot \frac{\text{ANPP}_M^{\text{LAI}_r}}{\text{ANPP}_C^{\text{LAI}_r}},$$

where the subscript $M$ denotes manipulation, $C$ denotes the control scenario and $\bar{X}$ indicates the annual scale. If the response ratios of $\beta$, LAI$_p$ and LAI$_r$ are independent at the annual scale, then

$$\frac{\overline{\text{ANPP}_M^{\text{CUE}}}}{\overline{\text{ANPP}_C^{\text{CUE}}}} \approx \frac{\overline{\text{ANPP}_M^{\text{GPP}_u}}}{\overline{\text{ANPP}_C^{\text{GPP}_u}}} \cdot \frac{\overline{\text{ANPP}_M^{\beta}}}{\overline{\text{ANPP}_C^{\beta}}} \cdot \frac{\overline{\text{ANPP}_M^{\text{LAI}_p}}}{\overline{\text{ANPP}_C^{\text{LAI}_p}}} \cdot \frac{\overline{\text{ANPP}_M^{\text{LAI}_r}}}{\overline{\text{ANPP}_C^{\text{LAI}_r}}} \cdot \frac{\overline{\text{ANPP}_M^{\beta}}}{\overline{\text{ANPP}_C^{\beta}}},$$

where overbars indicate average values for all years. This approximation is well-supported by the results of our simulations (Figure S6), even though data evidence suggests that CUE may change significantly under changes in water stress (Rowland et al., 2014). Using this decomposition in the model results, the average ANPP response ratio can be decomposed as the product of the average response ratios of $\beta$, LAI$_p$, LAI$_r$. Based on these considerations, we can attribute the changes in the modelled
ANPP among models to differences in simulated water stress, LAI dynamics, and phenological changes. Since only six (T&C, CABLE, JULES, TECO, DLEM and JSBACH) of the 10 participating models reported the water stress $\beta$ factor, this analysis was performed using this subset of models. All statistical analyses were performed in MATLAB 2019a.

3  |  RESULTS

3.1  |  Control scenario

Models captured the increasing trend of observed average ANPP to average $P$ across sites (Figure 1a). The RMSE between simulated and observed ANPP was in the range 23–354 g C m$^{-2}$ year$^{-1}$. Normalized RMSE of ANPP was weakly but positively correlated ($R^2 = .36; p = .067$) with the RMSE of normalized LAI (i.e. LAI divided by its maximum value). All models were positively biased. Positive biases can be partially attributed to model shortcomings but can be also explained by experimental underestimations in ANPP measurements (see Figure S1). Relative absolute biases, i.e. $\frac{\text{bias}}{\text{ANPP}_{\text{obs}}}$, are typically larger at the driest sites: $\frac{\text{bias}}{\text{ANPP}_{\text{obs}}}$ = $-6.3 \times 10^{-4}$ mm$^{-1}$ (estimated using ordinary least squares method; Table 2).

Both models and observations support a larger sensitivity of annual ANPP to interannual variation in precipitation at sites with intermediate wetness conditions (e.g. Garraf, Prades, Puèchabon, Konza; Figure 2). Specifically, in sites with a WI < 0.4 models (observations) have mean sensitivity $a_1 = 0.058(0.076)$ g C m$^{-2}$ mm$^{-1}$, in sites with $0.4 \leq WI < 1$ have $a_1 = 0.22(0.18)$ g C m$^{-2}$ mm$^{-1}$ and in sites with WI > 1

![Figure 1](a) Dependence of mean annual aboveground net primary production (ANPP) to average annual precipitation during the study period. Letters indicate observed values (L: Lahav; M: Matta; S: SGS; P: Prades; G: Garraf; K: Konza; Pb: Puèchabon; B: Brandbjerg; W: WB; Sb: Stubai). Lines indicate, for each model, a least square fit of a linear relationship: $\text{ANPP}(P) = \alpha P$ between the modelled mean annual ANPP and mean annual precipitation for all sites. (b) Standard deviation of modelled annual ANPP (circles) and observed annual ANPP (crosses) for all sites and models. Each model has a unique colour indicated in the legend [Colour figure can be viewed at wileyonlinelibrary.com]

### TABLE 2  Model skill across sites in terms of root mean square error (RMSE) for annual ANPP, normalized root mean square error (NRMSE) for annual ANPP, coefficient of determination for annual ANPP, average bias of ANPP, average bias of the standard deviation of ANPP, RMSE for daily LAI and RMSE for daily normalized LAI, i.e. $\frac{\text{LAI}_{\text{max}}(\text{LAI})}{\text{ANPP}_{\text{obs}}}$. 

<table>
<thead>
<tr>
<th>Model</th>
<th>ANPP − RMSE (g C m$^{-2}$ year$^{-1}$)</th>
<th>ANPP − normalized RMSE (−)</th>
<th>ANPP − $R^2$ (−)</th>
<th>ANPP − bias (g C m$^{-2}$ year$^{-1}$)</th>
<th>$\sigma(\text{ANPP})$ − bias (g C m$^{-2}$ year$^{-1}$)</th>
<th>LAI − RMSE (m$^2$)</th>
<th>LAI normalized RMSE (−)</th>
</tr>
</thead>
<tbody>
<tr>
<td>TC</td>
<td>76.318</td>
<td>0.368</td>
<td>0.8295</td>
<td>30.7907</td>
<td>$-13.5738$</td>
<td>1.2399</td>
<td>0.2956</td>
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<tr>
<td>JSBACH</td>
<td>233.0982</td>
<td>1.1239</td>
<td>0.2379</td>
<td>79.3713</td>
<td>$-19.6096$</td>
<td>1.2972</td>
<td>0.4276</td>
</tr>
<tr>
<td>DLEM</td>
<td>202.8963</td>
<td>0.9783</td>
<td>0.7732</td>
<td>96.935</td>
<td>$-23.7873$</td>
<td>1.2038</td>
<td>0.356</td>
</tr>
<tr>
<td>ORC MICT</td>
<td>121.7962</td>
<td>0.5872</td>
<td>0.6131</td>
<td>51.5792</td>
<td>$-5.7495$</td>
<td>1.1895</td>
<td>0.3966</td>
</tr>
<tr>
<td>ORC CNP</td>
<td>210.5444</td>
<td>1.0151</td>
<td>0.041</td>
<td>15.0756</td>
<td>$-1.0366$</td>
<td>1.1451</td>
<td>0.4198</td>
</tr>
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<td>ORCHIDEE</td>
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<td>0.549</td>
<td>0.6489</td>
<td>44.8288</td>
<td>9.3944</td>
<td>1.2675</td>
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<td>1.0399</td>
<td>0.4728</td>
<td>115.9473</td>
<td>$-5.1951$</td>
<td>2.147</td>
<td>0.3437</td>
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<td>JULES</td>
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<td>278.4353</td>
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<tr>
<td>TECO</td>
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<td>0.1123</td>
<td>0.982</td>
<td>5.3858</td>
<td>$-9.3174$</td>
<td>1.1347</td>
<td>0.3462</td>
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<tr>
<td>LPX</td>
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<td>0.5956</td>
<td>36.4618</td>
<td>33.2501</td>
<td>1.3886</td>
<td>0.4317</td>
</tr>
</tbody>
</table>

Abbreviations: ANPP, aboveground net primary production; LAI, leaf area index.
have $\bar{a}_1 = 0.13(-0.013) \text{ g C m}^{-2} \text{ mm}^{-1}$. At the most arid sites, annual precipitation explains a large fraction of the observed and modelled variability of annual ANPP, but the sites are not highly productive (i.e. absolute productivity values are low; Figure 1), yielding a low average sensitivity $a_1$. At the opposite end, mesic sites have higher productivity, but they are not water-limited during the observation period, resulting also in a low modelled sensitivity $\bar{a}_1$. Modelled sensitivity uncertainty was largest for intermediate precipitation regimes due to a larger model disagreement for those sites. For sites with a WI < 0.4, the average uncertainty, quantified here as the standard deviation of the distribution of simulated sensitivities $\sigma_{a_1 \text{inter}} = 0.24 \text{ g C m}^{-2} \text{ mm}^{-1}$ and for wet sites $\sigma_{a_1 \text{wet}} = 0.14 \text{ g C m}^{-2} \text{ mm}^{-1}$.

On average, the modelled sensitivity of ANPP to precipitation within sites was lower ($-0.15 \text{ g C m}^{-2} \text{ mm}^{-1}$) than ($-0.37 \text{ g C m}^{-2} \text{ mm}^{-1}$), estimated as the average slope of the linear models reported in Figure 1a) between sites. However, the uncertainty of the estimated temporal sensitivity from observations, as quantified by the 90% confidence limits of the linear model, is very high in most sites ($0.29 \text{ g C m}^{-2} \text{ mm}^{-1}$, averaged across all sites) and comparable to the uncertainty between models ($\sigma_{a_1} = 0.4 \text{ g C m}^{-2} \text{ mm}^{-1}$, averaged across all sites). A large uncertainty is related to either a small sample size, or low skill of the linear model. As a result, it is not possible to robustly quantify whether the modelled temporal sensitivities are statistically different from the observed ones, but overall only six of the 10 sites had mean modelled that were not non-statistically different than the one observed (Figure 2).

Simulated daily ET for the control simulations was substantially different regarding its day-to-day variability from measured ET at all three eddy sites (Konza, Puèchabon and WB). Correlation coefficients were in the range 0.27–0.78 with an average value between all models and sites of ~0.60 ± 0.13 (mean ± SD; Figure 3). Simulated variability of ET, expressed in terms of standard deviation at the daily scale, deviated substantially from the measured variability of

**FIGURE 2** Simulated and observed sensitivity of annual aboveground net primary production (ANPP) to annual precipitation ($a_1 = \partial \text{ANPP}/\partial P$). For each site, boxplots indicate the distribution of the simulated sensitivity of ANPP to precipitation by all models. Error bars show the sensitivity of observed ANPP to precipitation (blue squares) and the corresponding 90% confidence intervals (bar length) of the fit of the linear model. Crosses indicate the sites for which the mean value of the distribution of simulated sensitivities is not statistically different from the observed with 90% confidence. Sites are ranked from left to right in order of ascending wetness [Colour figure can be viewed at wileyonlinelibrary.com]

**FIGURE 3** Taylor diagrams for daily evapotranspiration (ET) and gross primary productivity (GPP) for all models and all sites with available flux tower data. Models are indicated with different colours according to the legend. Each site has a different marker (diamond for Konza, circle for WB and square for Puèchabon). The ideal model (i.e. reproducing precisely the data) would lie on the black markers, each corresponding to different sites [Colour figure can be viewed at wileyonlinelibrary.com]
ET. In particular, simulated variability from most models was lower than observed at Konza (observed $\sigma_{\text{ET}} = 1.76 \text{ mm/day}$, modelled $\sigma_{\text{ET}} = 1.40 \pm 0.3 \text{ mm/day}$), and higher than observed at Puèchabon (observed $\sigma_{\text{ET}} = 0.61 \text{ mm/day}$, modelled $\sigma_{\text{ET}} = 1.86 \pm 0.50 \text{ mm/day}$). For WB, the modelled ET variability was higher than observed, and inter-model agreement was low (observed $\sigma_{\text{ET}} = 1.39 \text{ mm/day}$, modelled $\sigma_{\text{ET}} = 1.51 \pm 0.45 \text{ mm/day}$). Seasonality of ET was well-reproduced by all models (Figure S2), partially explaining the high correlation coefficients (Figure 3). One pronounced exception is in Puèchabon, where the observed late summer reduction of ET and increase in early fall was reproduced only by a small subset of models (Figure S2).

Simulated daily GPP had a correlation ($-0.59 \pm 0.17$) with observed daily GPP for all models (Figure 3). A large fraction of the GPP correlation can be attributed to seasonality. However, the modelled variability was significantly different from the observed for all sites. Most models underestimated the daily variation of GPP at Konza (observed $\sigma_{\text{GPP}} = 4.04 \text{ g C m}^{-2} \text{ day}^{-1}$, modelled $\sigma_{\text{GPP}} = 2.87 \pm 1.88 \text{ g C m}^{-2} \text{ day}^{-1}$) and WB (observed $\sigma_{\text{GPP}} = 4.53 \text{ g C m}^{-2} \text{ day}^{-1}$, modelled $\sigma_{\text{GPP}} = 4.01 \pm 1.26 \text{ g C m}^{-2} \text{ day}^{-1}$) and overestimated the variability of daily GPP at Puèchabon (observed $\sigma_{\text{GPP}} = 1.68 \text{ g C m}^{-2} \text{ day}^{-1}$, modelled $\sigma_{\text{GPP}} = 2.67 \pm 1.01 \text{ g C m}^{-2} \text{ day}^{-1}$; Figure 3). Large model differences between observed and simulated GPP can be partially attributed to an incorrect representation of the magnitude of LAI. There is, indeed, a large disagreement between the modelled LAI across models (Figure 4). Modelled LAI is also significantly different than observed, even though LAI derived via remote sensing is also uncertain (Fang et al., 2013).

Model agreement in terms of ET and GPP varies also with timescale (Figure 5). In the driest sites (e.g. Lahav, Matta, SGS; WI < 0.4), models agree mostly with each other on the interannual variability of ET (average corr. coef. $\rho$ for ET at the annual ($y$) scale $\rho_{\text{ET, annual}} = 0.75$; for $\rho_{\text{GPP, annual}} = 0.35$). This is expected since at those sites annual ET almost equals the total amount of rainfall. However, a significant model
disagreement occurs at the daily \((d)\) scale \((\mu_{ET\_dry} = 0.58, \mu_{GPP\_dry} = 0.30)\). The opposite picture occurs in mesic sites \((WI > 1)\), where models agree better at the daily timescale for ET \((\mu_{ET\_wet} = 0.79)\), but their agreement is significantly lower at the annual scale \((\mu_{ET\_wet} = 0.61)\).

A similar pattern is also valid for GPP \((\mu_{GPP\_wet} = 0.77, \mu_{GPP\_wet} = 0.60)\); Figure 5).

Model agreement with regard to the dependence of the water stress factor \(\beta\) on root averaged soil moisture \((\theta(Z))\) is also low (Figure 6).

On average, model agreement was highest for sites with a large percentage of sand (Brandbjerg 88%–95% sand, Prades 48% sand) and lowest in sites with soils rich in more fine material (e.g. Lahav 22% sand, Matta 19% sand, SGS 14% sand, Konza 10% sand).

### 3.2 Manipulation experiments

Models were tested for their skill at reproducing changes in ANPP due to rainfall manipulations (Figure 6). Most models (75% for model-site-treatment combinations) correctly predicted the sign of the change in ANPP. However only 54% of the models for the drought treatment (10 models × 8 sites) and 43% for the irrigation treatment (10 models × 4 sites) have a mean response that is statistically similar in magnitude with the observed, highlighting a better model performance for rainfall exclusion than addition. The worst performance of the models was obtained for both the drought and irrigation experiments in Lahav and in the irrigation experiment in Konza where almost no model was able to capture the correct magnitude of the response ratio.

Even though observed ANPP estimated from biomass should be close to modelled ANPP (Figure S1), several uncertainties related to observations, such as the choice of biomass harvest date, the use of specific allometric equations and specific local conditions, could affect our results. For instance, the observed response to irrigation in Lahav and Matta is considerably different despite the two sites having similar vegetation and climate. Those differences are either due to measurement uncertainties, or due the large effect of some local properties (e.g. soil composition, nutrient availability; Golodets et al., 2013, 2015) causing significant changes in the ecosystem dynamics. Overall, the magnitude of responses is similar among models except CABLE, JULES and TECO, which show a larger sensitivity of ANPP to rainfall manipulation.
Modelled interannual variability of the responses was in most cases similar in magnitude to the observed for the rainfall exclusion experiments, and lower for the irrigation experiments (for the drought experiments, average modelled standard deviation of the response ratios was \( \sigma_{R_D} = 0.18 \); and observed \( \sigma_{R_{oD}} = 0.178 \). For irrigation experiments modelled standard deviation was \( \sigma_{R_I} = 0.25 \); and observed \( \sigma_{R_{oI}} = 0.42 \)). Outliers with regard to both the magnitude and the interannual variability of response ratios occurred for the most water-limited sites (Figure 7).

Besides carbon assimilation, changes in rainfall can simultaneously modify ET and thus the land surface energy balance. The coupling between ET and GPP depends heavily on the parametrizations of water stress and how this affects stomatal conductance and the reduction of photosynthesis. It further depends on vegetation dynamics such as a shift of carbon allocation from leaves to roots or leaf shedding due to water stress. To quantify the responses of the ET and GPP coupling, we compute the relative changes of WUE for the various cases (Figure 8). Most models predict relatively small changes in WUE (i.e. \( R \approx 1 \)) for both drought (\( R_{oD}^m = 0.98 \)) and irrigation (\( R_{oI}^m = 1.08 \)) treatments, indicating a change of comparable magnitudes for both ET and GPP. CABLE, JULES and TECO occasionally simulate larger changes, in both positive and negative directions, in WUE for the most water-limited sites. This larger change can be attributed to a more sensitive response of GPP to water stress than ET.

### 3.3 Response attribution

We partitioned the total response ratio of ANPP into relative changes of (a) the \( \beta \) stress factor; (b) peak LAI (LAI\(_p\)); and (c) the length of the growing season approximated by LAI\(_r\) (Figure 9). Changes in simulated ANPP following rainfall manipulation can be almost exclusively attributed to changes in \( \beta \) and LAI\(_p\). The response ratio of LAI\(_r\) was always close to unity \( R_{LAI_r} = 0.98 \pm 0.058 \) (mean ± SD) for the drought
treatment and $R_{\text{LAI}} = 1.01 \pm 0.029$ for the irrigation treatment contributing insignificantly to the response ratio of ANPP. Thus, no model predicted substantial changes in the length of the growing season. A reduction or enhancement of $\beta$ for the drought and irrigation experiments explained the largest fraction of ANPP responses at wet sites, but the uncertainty of the relative strengths of changes in $\beta$ and LAI$p$ was high (drought treatment for sites with WI > 1, $R_\beta = 0.95 \pm 0.08$, $R_{\text{LAI}} = 0.91 \pm 0.18$; irrigation treatment for sites with WI > 1, $R_\beta = 1.05 \pm 0.06$, $R_{\text{LAI}} = 1.02 \pm 0.02$). For the driest sites both LAI$p$ and $\beta$ explained a large fraction of the total response for the drought treatment, whereas LAI$p$ was the dominant and simultaneously the most uncertain factor for the irrigation treatment (drought treatment for sites with WI < 0.4, $R_\beta = 0.87 \pm 0.10$, $R_{\text{LAI}} = 0.77 \pm 0.24$; irrigation treatment for dry sites with WI < 0.4, $R_\beta = 1.06 \pm 0.10$, $R_{\text{LAI}} = 1.49 \pm 0.86$). Differences in the simulated responses of both $\beta$ and LAI$p$ among models were high as indicated by the standard deviations above. At the sites where rainfall exclusion was applied only in the irrigation treatment for dry sites with WI < 0.4, LAI$p$ was the dominant and simultaneously the most uncertain factor for the irrigation treatment (drought treatment for sites with WI > 1, $R_\beta = 0.93 \pm 0.09$, $R_{\text{LAI}} = 0.78 \pm 0.27$), but given the large variability among models, it is not possible to conclude if this is a true signal. The variability was higher for the most water-stressed sites, primarily because for those sites model disagreement on the estimated response ratio of ANPP was also the highest.

4 | DISCUSSION

4.1 | Multisite and local sensitivities to rainfall and the role of temporal scales

Most models overestimated the relationship between mean annual precipitation and average annual ANPP observed across sites, but managed to capture well the overall trend, despite large site differences in terms of vegetation coverage and overall climatic regime (Figure 1). This result confirms that terrestrial biosphere models can capture spatial gradients of vegetation productivity relatively well (e.g. Wu et al., 2018). Reproducing local (single-site) response of ANPP to interannual precipitation variability has been generally found to be more challenging (Fatichi & Ivanov, 2014). In fact, previous intercomparison studies have found that models have significant biases at various timescales, from subdaily (Matheny et al., 2014) to decadal (Dietze et al., 2011). Dietze et al. (2011) found model errors to be largest at the annual scale. In agreement with such a result in our experiment, models differed greatly in their simulated sensitivity of local-scale productivity to annual precipitation but were able to reproduce the previously reported stronger spatial than temporal sensitivity of productivity to rainfall. A large model disagreement with regard to the magnitude of the interannual variability of ANPP also confirms the previously found difficulties of models to properly capture carbon dynamics at the annual scale (e.g. Dietze et al., 2011; Paschalis, Fatichi, Katul, & Ivanov, 2015). Despite large model disagreement we found that the within-site sensitivity of ANPP to precipitation is lower than across-site sensitivity of ANPP to average precipitation, in agreement with a number of previous observational (Goward & Prince, 1995; Huxman et al., 2004; Knapp & Smith, 2001) and modelling results (Fatichi & Ivanov, 2014; Wu et al., 2018).

One of the main reasons for model disagreement originates from the differences in parametrization in schemes representing water limitation effects on water and carbon fluxes (e.g. Trugman et al., 2018), summarized here by the water stress parameter $\beta$ (Figure 6). Those parametrizations influence ecosystem dynamics at a wide range of temporal scales, complicating assessment of their skill. For instance, at shorter timescales (e.g. daily), in ecosystems with no water limitation, where temperature and radiation are the dominant controls for ET and GPP (Paschalis et al., 2015), models had a high agreement (Figure 5), in terms of correlation. This highlights that parametrizations that impact the temporal changes of ET and GPP should be relatively consistent among models, at least during wet conditions (Ukkola et al., 2016). Even though correlation between models was high, large variability between models with regard to the actual magnitude of the fluxes was pronounced (Figures S2–S4), primarily for carbon fluxes (e.g. GPP). This indicates that a ‘scaling’ factor affecting GPP is significantly different among models. For our experiments, LAI could be this explanatory ‘scaling’ factor (Figure 4), as models greatly differed regarding the seasonality and magnitude of LAI.

Significant changes emerge under drought, when water stress parametrizations influence the simulation of water and carbon fluxes. Different water stress parametrizations alter the water/carbon dynamics at different scales. In severely water-limited systems (WI < 0.4), model results diverge in terms of GPP and ET at short temporal scales (e.g. daily; Figure 5). Thus, parametrizations of how water stress impacts processes operating at daily and subdaily timescales are crucial, and highly diverging among models. Such parametrizations include stomatal regulations and downregulation of photosynthesis during drought. In general, plant hydraulic dynamics will also operate at these temporal scales, but none of the participating models simulated such processes in detail. In severely water-limited ecosystems, the amount of annual precipitation imposes a strong constraint on ET (i.e. ET $\propto$ P), leading to overall good agreement between models for annual ET. However, this agreement is not true for transpiration alone (Figure S8), highlighting the major importance of how stomatal limitations are implemented in models. Physical constraints for productivity are not as strong, and thus models have large disagreement with respect to GPP even at annual scales.

In intermediate wetness sites (0.4 $\leq$ WI $< 1$), in our simulations, models disagree at intermediate scales (weeks–months) in terms of GPP (consistent with the wavelet coherence analysis presented at Figure S7). As mentioned before, at short (daily) temporal scales, temperature and radiation mostly determine water and carbon fluxes, when water is not a strong limiting factor, and due to the similar parametrizations among models (Wu et al., 2018), we detect a substantial convergence in GPP. However, since such controls ‘fade’ with increasing temporal scales, the effects of features linked to soil moisture dynamics, such as the soil moisture retention after a rainfall event, can manifest at longer temporal scales (Paschalis et al., 2015). Those dynamics can be influenced by factors including both
biotic and abiotic factors such as the parameterizations of soil properties that determine the temporal dynamics of soil moisture and the vertical distribution of root biomass, affecting how plants withdraw water from the soil. In fact, models were found to strongly disagree on how plants are affected by soil moisture (biotic factor—Figure 6) and on the soils’ water holding capacity, as indicated by the range of accessible values of soil moisture (abiotic factor—Figure 6).

At the wettest sites (WI > 1), strong model disagreement in terms of both water and carbon fluxes occurs at annual scales. A key factor for model disagreement for those sites is LAI (Figure 4). Model disagreement in LAI is a composite effect of the water stress impacts to LAI development and the overall model disagreement in leaf pholology and carbon allocation rules (Figure 4; Richardson et al., 2012).

All those behaviours highlight further the need to correctly capture water/carbon dynamics at multiple timescales, from the scale of the individual rain pulse (Huxman et al., 2004) up to interannual scales where drought legacies can have an important effect (Anderegg et al., 2015). The need to understand in detail multiscale dynamics linked to water stress and soil moisture dynamics is also exacerbated by the fact that model disagreement in terms of the sensitivity of ANPP to annual rainfall is highest for sites with intermediate wetness (0.4 ≤ WI < 1). Those regions experience moderate water limitations, and the impact of water limitation to fast-acting processes (changes in e.g. stomatal conductance, photosynthesis) can accumulate and impact longer timescales through slow-acting processes (e.g. changes in LAI). Additionally, areas with intermediate wetness are expected to operate close to soil moisture thresholds inducing plant water stress. Sensitivity of the responses of ANPP to precipitation in those sites is concurrently the highest and most uncertain (Figure 2). This can have a large impact on our ability to model the fate of terrestrial CO₂, given that those areas are among the largest contributors to the interannual dynamics of the growth rate of CO₂ (Ahlström et al., 2015; Poulter et al., 2014). Understanding such dynamics across scales requires high quality and high frequency long-term measurements, not only for CO₂ and water fluxes but also for soil moisture dynamics (Vicca et al., 2012). Annual ANPP values alone are limiting our inference capabilities and even 10–20 years of annual ANPP data were not long enough to obtain a precise estimate of the sensitivity of ANPP to precipitation.

Uncertainties arise from the relatively short span of the record, but also due to the lack of data describing short-scale dynamics of carbon assimilation and growth in manipulation experiments. Annual precipitation has been found to be a relatively weak descriptor of the interannual variability of water and carbon fluxes in many locations worldwide (Fatichi & Ivanov, 2014). A better descriptor would be the time duration during a year when favourable meteorological conditions for photosynthesis occur under well-watered conditions (Fatichi & Ivanov, 2014; Zscheischler et al., 2016). As a result, a few bursts of positive extremes in terms of productivity can strongly modify the annual budget and long-term dynamics (Zscheischler, Mahecha, et al., 2014). Therefore, to quantify the interannual dynamics of vegetation productivity, detailed knowledge of water/carbon fluxes, meteorology, soil moisture and plant water status at fine-temporal scales would be essential. In fact, previous research at the PHACE experiment, one of the few facilities that combined such high frequency measurement clearly identified the problems models have in reproducing sub-annual dynamics (De Kauwe et al., 2017). Given the present limited availability of such data, new ways of combining existing data (e.g. combining different data-streams representing short and long-term-dynamics in multiple locations, such as Fluxnet sites for water and carbon fluxes at high frequencies, sites equipped with phonecams for high-frequency phenology monitoring, soil moisture networks (e.g. COSMOS, the International Soil Moisture Network, the Long Term Ecological Research Network, etc.), open access data archiving with common data formats to facilitate data exchange between research groups and the use of proxy data to extend the length of the time series (e.g., tree rings) are necessary to better inform models (Babst et al., 2018; Pappas, Mahecha, Frank, Babst, & Koutsoyiannis, 2017).

### 4.2 Response to manipulation experiments

The modelled sensitivities of vegetation dynamics to changes in rainfall are highly uncertain. On average, most models captured better the observed responses of vegetation to rainfall exclusion than addition (Figure 7). That behaviour can be associated with low skill in reproducing the asymmetric response of productivity to precipitation (Wu et al., 2018), failing to capture the correct pattern of the productivity saturation effect associated with rainfall increase.

Even though, multiple models generated close vegetation productivity responses in the rainfall exclusion experiments, the underlying reasons are very different and at the same time highly uncertain (Figure 9). In the more water-limited ecosystems, both changes in LAI magnitude and the level of plan water limitation determine productivity responses. Variability of the relative strength of β and LAI between models is large. Variability concerning LAI is larger than β, which can be explained by the fact that LAI integrates the model differences related to LAI phenology, carbon allocation rules and reductions in photosynthetic rates due to soil moisture limitations. Pinpointing which model best captures the relative strengths of changes in β and LAI would require simultaneous high-frequency data, including soil moisture, regular measurements of stomatal conductance and leaf water potentials, high-frequency photosynthetic rates and regular LAI estimates. At more mesic sites, physiological effects of water stress (through β) are the main reason for productivity responses. The reason is that in such sites, induced water stress is mild. Productivity will be reduced during the imposed water stress due to rainfall exclusion, but this small increase in water stress cannot cause large changes in vegetation structure (Estiarte et al., 2016), or LAI.

Disagreement in irrigation experiments is primarily related to leaf area dynamics. The reason can be that in the simulations where water stress was relieved, model disagreement originates primarily from the leaf area dynamics simulated for the unstressed
conditions. Those dynamics are related to the choice of carbon allocation and leaf phenology algorithms. Pronounced model differences related to those dynamics can be shown via the magnitude and seasonal patterns of LAI (Figure 4) as simulated by all models. Both the allocation and the phenology algorithms affect the dynamics of LAI. In our simulations (Figure 4) the range in modelled LAI is large and comparable with that reported by previous studies (De Kauwe et al., 2017; Walker et al., 2014). Parametrizations of carbon allocation rules are also limited by the use of generic PFTs used by most models. Such a choice is generally very restrictive and cannot capture the natural variability of plant traits, which is relevant at the local scale.

In our analysis changes in growing season length were not evident and did not influence out results. This is not surprising, as all rainfall manipulation experiments decreased or increased the available water to the ecosystem, without altering its 'pulse' structure, including the frequency of rainfall occurrence, and the time of storm arrival (Ross et al., 2012). As vegetation phenology in water-limited ecosystem is very sensitive to the pulse structure dynamics of rainfall (Heisler-While et al., 2009), evaluating in future experiments, whether models can properly capture the responses of vegetation to rainfall pulses in terms of productivity and drought deciduousness is very important. Changes in rainfall pulses will also strongly impact soil respiration dynamics that will contribute significantly to the total carbon balance (Jarvis et al., 2007; Unger, Máguaes, Pereira, David, & Werner, 2010).

4.3 | Outlook for model developments and observations

Our results highlight the need for a coordinated effort of new model development and data collection that could enable validations that are much more detailed than currently achievable here. Model discrepancies in the present study were attributed to the $\beta$ stress factor and long-term leaf area dynamics. The models used in this study implemented simple conceptual, yet vastly different (Wu et al., 2018) parametrizations of the effects of water limitation, neglecting plant hydraulics and thus impacts on the water transport system (xylem cavitation) that can lead to hydraulic failure and/or carbon starvation (Bonan et al., 2014; McDowell, 2011; McDowell et al., 2013; Xu, Medvigy, Powers, Becknell, & Guan, 2016). This could be an important limitation. However, tree mortality is not a prominent feature of the manipulation experiments considered here and while it has attracted a lot of attention, models first need to better simulate mild to severe water stress before considering vegetation death. For instance, differences associated with the $\beta$ factor are not only related to plant physiological thresholds but are associated with a complex function of the assumed soil textural properties. Those properties are translated into soil hydraulic parameters (Van Looy et al., 2017), affecting soil moisture dynamics and ET and ultimately their interplay with the value of the $\beta$ factor. It is currently impossible or very difficult to identify which model is more realistic in this respect and each model can only 'tune' all the above components at once. Specialized experiments measuring for example simultaneously high-frequency water and carbon fluxes, soil moisture and plant water status in controlled environments could be designed to develop more informed parameterizations of $\beta$, and eventually expand to more detailed mechanistic representation of ecosystem-scale plant hydraulics (Anderegg et al., 2016; Konings & Gentine, 2017).

Correct modelling of leaf area dynamics is equally important as the plant physiological stress $\beta$ for quantifying the effect of rainfall changes in ecosystem functioning (Yang, Medlyn, De Kauwe, & Duursma, 2018). Simulation of LAI could be constrained better than currently done with available information, considering that high-frequency LAI measurements in an experiment could be added with a relatively low budget. Observations of LAI, via indirect methods, are common at large scale. Extensive ground (Iio, Hikosaka, Anten, Nakagawa, & Ito, 2014) and remote-sensing estimates (Zhu et al., 2013) of LAI and phenology data from low-cost cameras worldwide (Brown et al., 2016; Klosterman et al., 2014) can be used to further constrain phenology and carbon allocation. Regarding carbon allocation, belowground dynamics and their responses to water limitation should also be simultaneously quantified.

From an observational perspective, in order to improve models, we need to disentangle the effects on plant physiological stress from those on vegetation dynamics at the local scales. Since physiological effects of water stress manifest earlier than changes of LAI or carbon pools, a nearly continuous monitoring of photosynthesis, ET, leaf and soil water potentials, sap flow and LAI would be essential to get further insights. These quantities are often observed (e.g. using eddy covariance systems, sap flow sensors, leaf porometers, hyperspectral cameras), but rarely in an integrated manner and associated with rainfall manipulation experiments. This should become a priority to foster model developments.

Finally, new streams of data via remote sensing can be also used for detailed model confirmation at larger scales. Satellite and airborne data related to vegetation structure, spanning from leaf chemistry to delineation of individual trees (Andersen, Reutebuch, & McGaughey, 2006; Asner & Martin, 2009; Gougeon & Leckie, 2006; Vicca et al., 2016), high frequency photosynthesis through solar induced fluorescence, soil moisture (Liu et al., 2011), and plant hydraulic status (Konings & Gentine, 2017) currently exist. Such data can help us to identify the mechanistic link between plant water stress and how it affects vegetation productivity from short-term photosynthesis reduction to decadal scales involving plant mortality and composition shifts. Note however that estimates of photosynthetic activity during water stress purely based on remote sensing (light reflection signals) are often biased and need to be interpreted with care (De Kauwe, Keenan, Medlyn, Prentice, & Terrer, 2016; Stocker et al., 2019).

In conclusion, our key finding in this study is that current generation terrestrial biosphere models have major uncertainties related to simulating plant water stress, and its impact on the terrestrial carbon cycling. Those uncertainties arise from the model formulations related
to both carbon allocation patterns and phenology and the representation of water stress frequency and magnitude on carbon assimilation. These two effects are inherently coupled at a wide range of scales. To decouple the two effects and constrain mechanistic representations of how water stress acts on multiple processes will require the close collaboration between experimentalists and modellers, for planning and implementing new ‘high frequency’ experiments (Rineau et al., 2019). These experiments should observe across a range of temporal scales from hourly values of photosynthesis and ET, to daily and weekly LAI dynamics, up to arrive to annual changes in species composition (Halbritter et al., 2019).

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CONFLICT OF INTEREST
The authors declare no conflict of interest.

DATA AVAILABILITY STATEMENT
All meteorological input data and model outputs can be found at the zenodo data repository (https://zenodo.org; doi: https://doi.org/10.5281/zenodo.3627959).

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REFERENCES


Estiarte, M., Vicca, S., Peñuelas, J., Bahn, M., Beier, C., Emmett, B. A., ... Janssens, I. A. (2016). Few multiplier precipitation-runoff experiments find a shift in the productivity-precipitation relationship.


are an important control of interannual variability in carbon and water fluxes. *Journal of Geophysical Research: Biogeosciences*, 121(8), 2186–2198. https://doi.org/10.1002/2016JG003503


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